



# LEARNING DYNAMIC SYSTEMS: MARKOV MODELS

Markov Process and Markov Chains  
Hidden Markov Models  
Kalman Filters



# Types of dynamic systems

		Problem of future state prediction
<b>Predictability</b>	<b>Easily predictable state</b>	<b>Hardly predictable state: state noise</b>
<b>Observability</b>		
<b>Easily observable state</b>	Trajectory of a satellite	Trajectory of a GPS localized vehicle
<b>Hardly observable state: measurement noise</b>	Controlled underwater vehicle with inertial navigation system in calm water	radiolocalized phones using GSM cells triangulation
<b>Partially observable state + measurement noise</b>	Trajectory of an indoor controled robot (SLAM)	Phone call localization, speech recognition
		Problem of current state estimation (filtering) Problem of past state estimation (smoothing) Problem of past state trajectory

# Markov models: a global view

<i>Observability Type of State</i>	<i>Observable state</i>	<i>Partially observable state</i>
<i>Discrete</i>	<b>Markov Chains</b>	<b>Hidden Markov models (HMM)</b>
<i>Linear continuous models</i>	<b>Kalman filter, ARMA models, etc</b>	
<i>Non-linear continuous models</i>	Extended Kalman filter, Particle filters, etc	



# LEARNING DYNAMIC SYSTEMS: MARKOV MODELS

**Markov Process and Markov Chains**

Hidden Markov Models

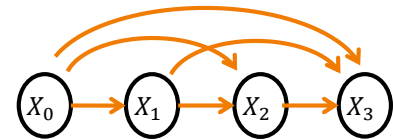
Bayesian Filtering and Kalman Filter



# Markov process

**Stochastic process:** sequence of random variables  $(X_0, \dots, X_t, \dots)$

$$P(X_0, \dots, X_t) = P(X_0) \prod_{i=1}^t P(X_i | X_0, \dots, X_{i-1})$$



**Markov process:**

« Knowing the past doesn't help to predict the future when the (close) present is known. »

$$P(X_0, \dots, X_t) = P(X_0) \prod_{i=1}^t P(X_i | X_{i-1}, \dots, X_{\max(i-k, 0)})$$

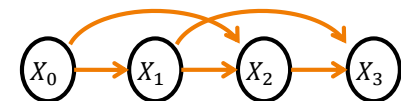
$X$  is the system **state**,  $k$  is the **order**.

**Examples :**

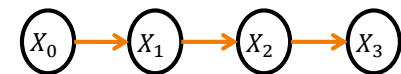
Linear autoregressive models (AR) for prediction in economy/finance

$$X_t = a_1 X_{t-1} + \dots + a_k X_{t-k} + \varepsilon_t$$

where  $\varepsilon_t \sim \mathcal{N}(0, \sigma_t^2)$  is a normal white noise



Markov process of order 2



Markov process (of order 1)

# Markov Chain

## Definition:

- Markov process of order 1
- with **observable discrete state** (from 1 to  $n$ )

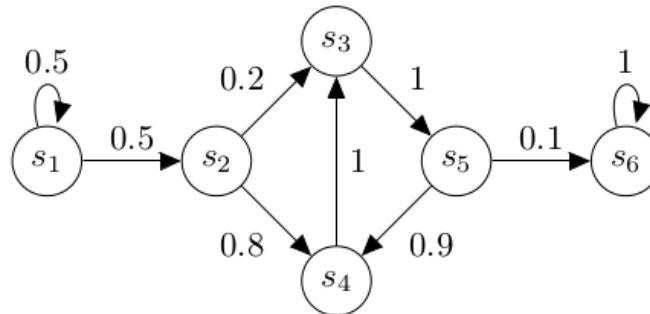
**Parameterization:** distributions of initial state and **transitions:**  $\theta = \left( (p_i^0)_i, (p_{i,j}^t)_{t,i,j} \right)$

$$\boxed{p_i^0 = P(X_0 = i)} \text{ and } \boxed{p_{i,j}^t = P(X_t = i | X_{t-1} = j)}$$

**Homogeneous chain:** time independent

$$\boxed{p_{i,j} = P(X_t = j | X_{t-1} = i)}$$

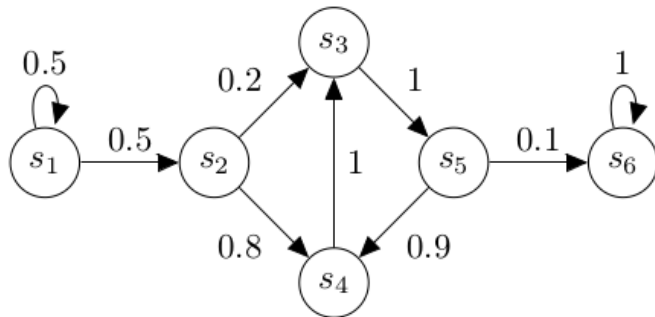
**Representation as a graph of a stochastic finite state machine:**



# Matrix representation of Markov Chains

A Markov chain with  $n$  state is defined by a  $n \times n$  **transition matrix**:

$$\mathbb{P}_t = (p_{i,j}^t) = (P(X_t = j | X_{t-1} = i))_{i,j}$$



$$\mathbb{P} = \begin{pmatrix} 0.5 & 0.5 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0.2 & 0.8 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0.9 & 0 & 0.1 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}$$

**Property:**  $\mathbb{P}_t$  is **stochastic**

$$\forall i, \sum_{j=1}^n p_{i,j}^t = 1$$

# State prediction from a Markov chain

**Fundamental property:** given state distribution  $P_t = \begin{pmatrix} P(X_t = 1) \\ \vdots \\ P(X_t = n) \end{pmatrix}$ ,

$$P(X_t = j) = \sum_{i=1}^n P(X_t = j | X_{t-1} = i) \cdot P(X_{t-1} = i) \Leftrightarrow \boxed{P_t = \mathbb{P}_t^T \times P_{t-1}}$$

**Predicting state at «  $t + k$  » :**

Oubli de la transposition  
dans le poly!!!

General case:  $P_{t+k} = \prod_{h=1}^k \mathbb{P}_{t+h}^T \times P_t$

Homegenous case:  $P_{t+k} = \mathbb{P}^{T^k} \times P_t$

**Likelihood of an observation**  $(x_0, \dots, x_T)$  :

$$P(x_0, \dots, x_T | \theta) = P(X_0 = x_0) \prod_{h=1}^T (\mathbb{P}_i)_{x_{i-1}, x_i}$$



# Learning a Markov Chain

Given  $n$  i.i.d state sequences:

$$s_1 = (x_0^1, \dots, x_{T-1}^1), \dots, s_n = (x_0^n, \dots, x_{T-1}^n)$$

**What is the underlying Markov Chain?**

MLE estimation  $\rightarrow$  frequencies:

$$\hat{p}_{ij}^t = \frac{|k \text{ s.t. } x_t^k = j \text{ and } x_{t-1}^k = i|}{|k \text{ s.t. } x_{t-1}^k = i|}$$

**Problem:** many coefficients are likely to be equal to  $\frac{0}{0}$

$\rightarrow$  Introduction of a Dirichlet prior:  $(p_{ij}^t)_{1 \leq j \leq n} \sim \text{Dir} \left( (\alpha_{ij}^t)_{1 \leq j \leq n} \right)$

$$\hat{p}_{ij}^t = \frac{|k \text{ s.t. } x_t^k = j \text{ and } x_{t-1}^k = i| + \alpha_{ij}^t}{|k \text{ s.t. } x_{t-1}^k = i| + \sum_j \alpha_{ij}^t}$$

In general  $\alpha_{ij}^t = 1$  (uniform distribution)

# Example of application:



One has a corpus of classical music scores with the name of their composer.

## 1) Problem of supervised classification:

Given a new score, guess the composer.

## 2) Problem of trajectory generation:

Generate a score that sounds like composed by two given composers



Xenakis and the  
« stochastic music »

# Solution:

## Learn a homogeneous Markov Chain of order $k$ for each composer $c$

- State space = (pitch, length) of a note
- Compute:

$$\hat{p}^c_{i_k, \dots, i_1, j} = \frac{|t \text{ s. t. } X_t = j \text{ and } X_{t-1} = i_1 \dots \text{ and } X_{t-k} = i_k| + 1}{|t \text{ s. t. } X_{t-1} = i_1 \dots \text{ and } X_{t-k} = i_k| + n}$$

- Find optimal  $k$  by cross validation

## Predict composer for score $(x_0, \dots, x_T)$ :

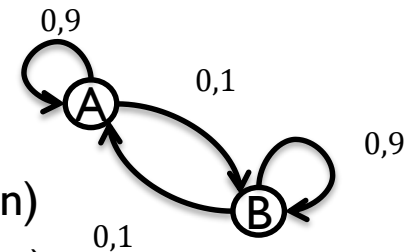
- For each composer, compute likelihood:

$$P(x_0, \dots, x_T | C = c) = p_c(x_0) \prod_{t=1}^T \hat{p}^c_{x_{t-k}, \dots, x_{t-1}, x_t}$$

- Choose  $\hat{c} = \underset{c}{\operatorname{argmax}}(P(x_0, \dots, x_T | C = c))$

## Problem of trajectory generation:

- Averaging both transition matrices is a bad idea (loss of information)
- Use a hierarchical Markov chain with two states (one per composer)
- Generate a note from the Markov chain of the current composer

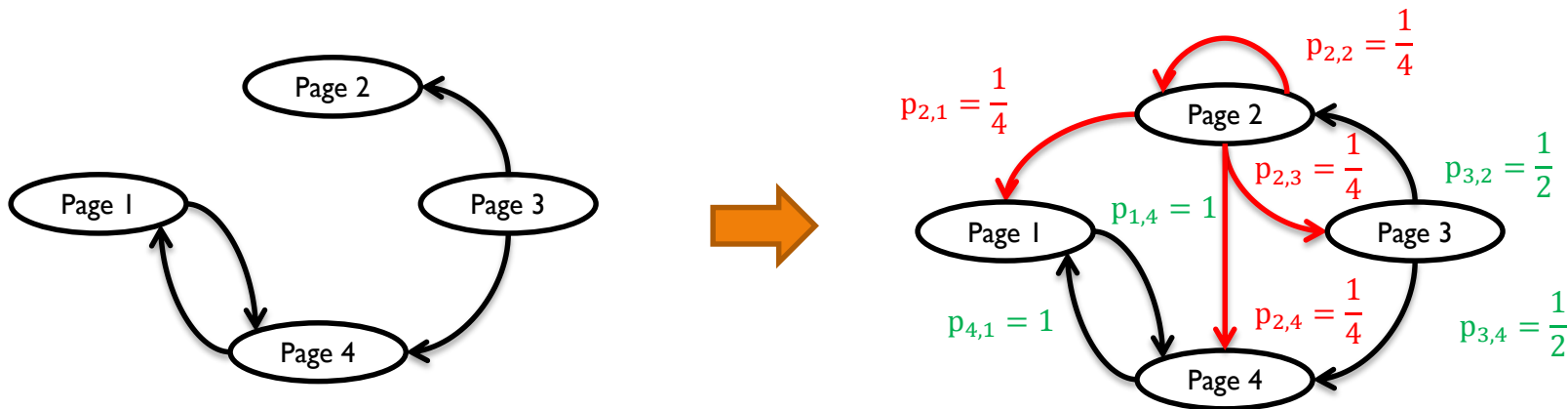


# Other application: PageRank at Google

Compute importance index for nodes in a network (Web, etc)

Let's  $X_t$  be the current page of a random Websurfer (random walk model)

Page rank value of page  $p$  = asymptotic probability  $\lim_{t \rightarrow \infty} P(X_t = p)$



**Model Web as a Markov Chain**  $p_{ij} = P(X_t = j | X_{t-1} = i)$

- If  $\text{deg}(i) \neq 0$ , if "i links to j"  $p_{ij} = \frac{1}{\text{deg}(i)}$  otherwise  $p_{ij} = 0$
- If  $\text{deg}(i) = 0$ ,  $p_{ij} = \frac{1}{n}$

**Does  $\lim_{t \rightarrow \infty} P(X_t = p)$  exist and is independent of  $X_0$ ?**

# Stationary distribution and equilibrium

## Equilibrium state distribution:

state distribution limit independent of the initial state distribution

$$\exists P_\infty, \forall P_0, \lim_{t \rightarrow \infty} P_t = \lim_{t \rightarrow \infty} (\mathbb{P}^{T^t} \times P_0) = P_\infty$$

**Stationary distribution** : state distribution  $P_\infty$  that is a fixed point

$$P_\infty = \mathbb{P}^T \times P_\infty$$

**Property 1**: an equilibrium distribution is stationary

**Property 2**: every Markov chain has some stationary distribution(s)

*Every stochastic matrix accepts 1 as the largest eigenvalue (in absolute value).*

*Components of right and left eigenvectors for eigenvalue 1 have all the same sign.*

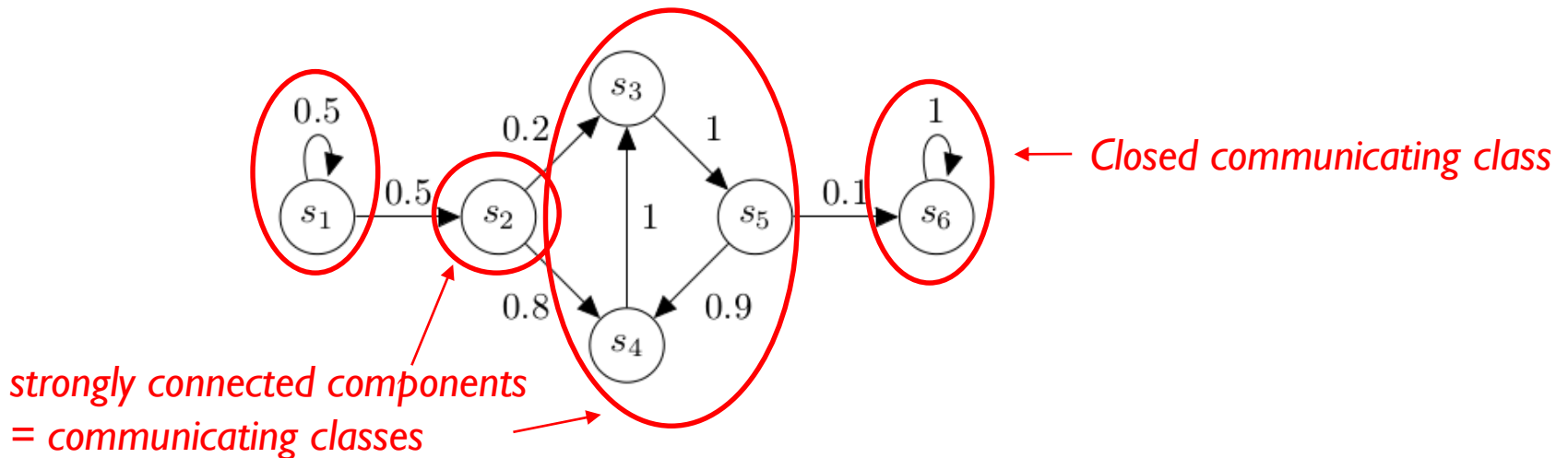
$$\mathbb{P}^T = \begin{bmatrix} 0.5 & 0 & 0 & 0 & 0 & 0 \\ 0.5 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0.2 & 0 & 1 & 0 & 0 \\ 0 & 0.8 & 0 & 0 & 0.9 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0.1 & 1 \end{bmatrix} \Rightarrow \Lambda = \begin{bmatrix} -0,48 - 0.83i \\ -0,48 + 0.83i \\ 0 \\ 0.5 \\ 0.97 \\ 1 \end{bmatrix} \Rightarrow P_\infty = \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 1 \end{bmatrix}$$

# Notion of reducibility

- State  $s'$  **accessible** from state  $s$  ( $s \rightarrow s'$ ) if

$$s \rightarrow s' \Leftrightarrow \exists t \geq 0, P(X_t = s' | X_0 = s) > 0$$

- $s$  and  $s'$  **communicate** ( $s \leftrightarrow s'$ ) if  $s \rightarrow s'$  and  $s' \rightarrow s$
- Communicating classes** are equivalent classes for  $\leftrightarrow$
- A **closed** communicating class has no outgoing link.
- Theorem:** the number of stationary distributions is the number of closed communicating classes
- A chain is **irreducible** if there is only one (closed) communicating class, i.e. the transition graph is **strongly connected**

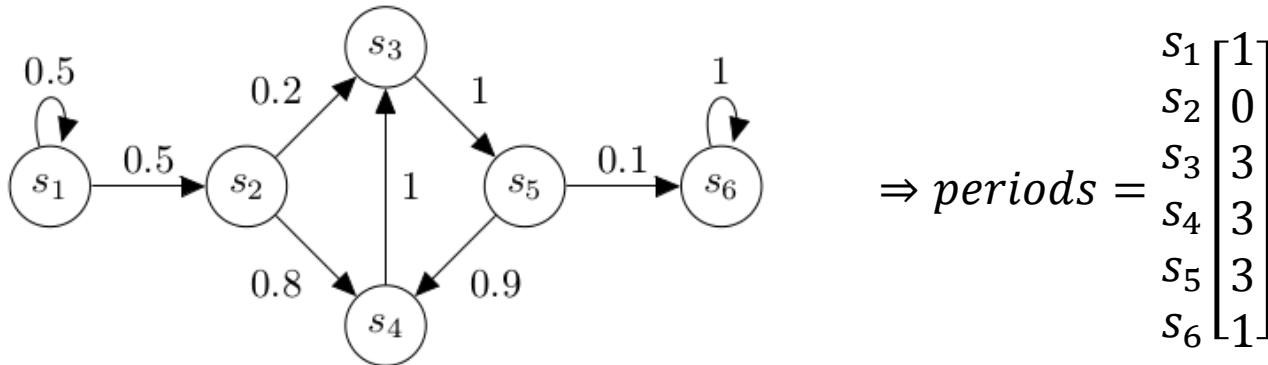


# Notion of periodicity

- **Period** of a state is

$$period(s) = \gcd(\{t | P(X_t = s | X_0 = s) > 0\})$$

- A Markov chain is **aperiodic** if all states have a period equal to 1.



## Theorem: sufficient condition for convergence to an equilibrium

An homogeneous **irreducible** and **aperiodic** Markov chain converges to an equilibrium

# Back to PageRank

**Problem:** the Markov chain of the Web is neither irreducible nor aperiodic

**Solution:** every **complete** transition graph is irreducible and aperiodic

$$\forall i, \forall j, p_{ij} > 0$$

**Algorithm:** for every page, one draws a number between 0 and 1

- If  $x \geq \alpha$ , chooses randomly an outgoing link
- Otherwise, teleport randomly to a page of the Web

**Consequence:** new transition graph is complete:

$$\forall i, \forall j, p'_{ij} = (1 - \alpha)p_{ij} + \alpha \cdot \frac{1}{n} > 0$$

$$\mathbb{P}^T = (1 - \alpha) \begin{bmatrix} 0.5 & 0 & 0 & 0 & 0 & 0 \\ 0.5 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0.2 & 0 & 1 & 0 & 0 \\ 0 & 0.8 & 0 & 0 & 0.9 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0.1 & 1 \end{bmatrix} + \frac{\alpha}{n} \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \times [1 \quad 1 \quad 1 \quad 1 \quad 1 \quad 1]$$





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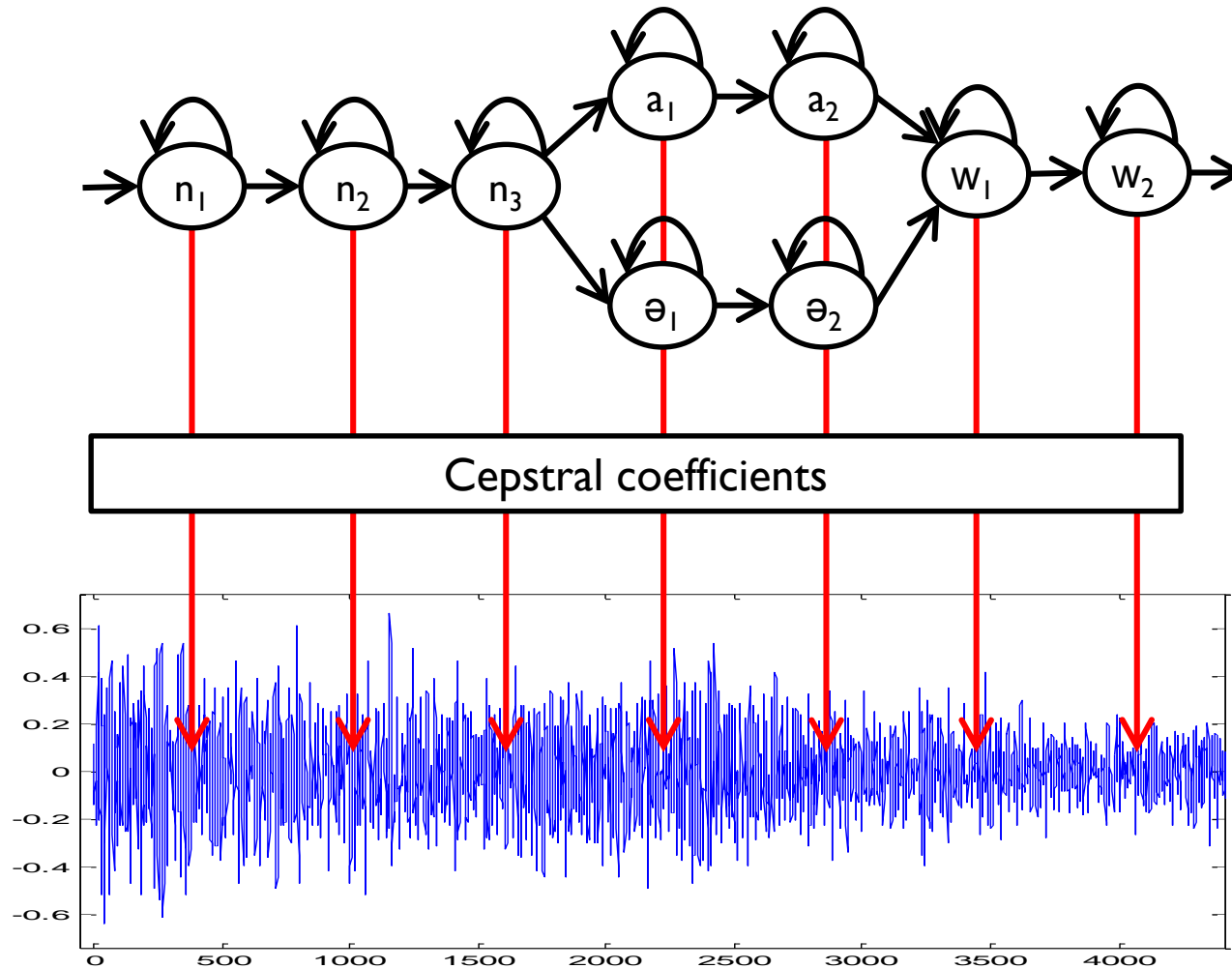
Markov Process and Markov Chains

**Hidden Markov Models**

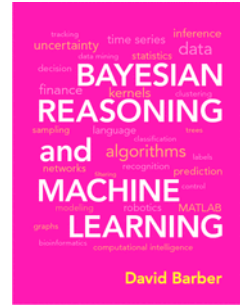
Bayesian Filtering and Kalman Filter



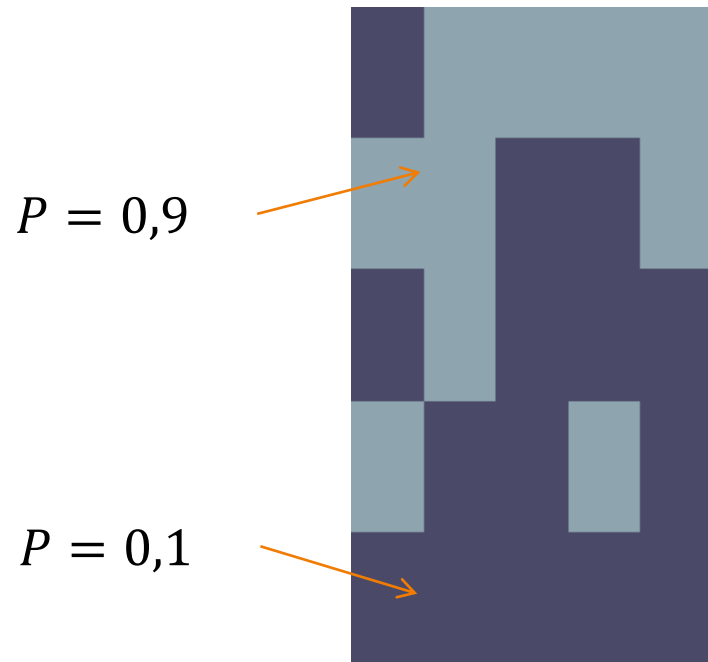
# Example of application: Speech recognition systems



# Example of a partially observable state: the burglar problem (Barber)



A burglar walks on a grid 5 x 5 in the dark.

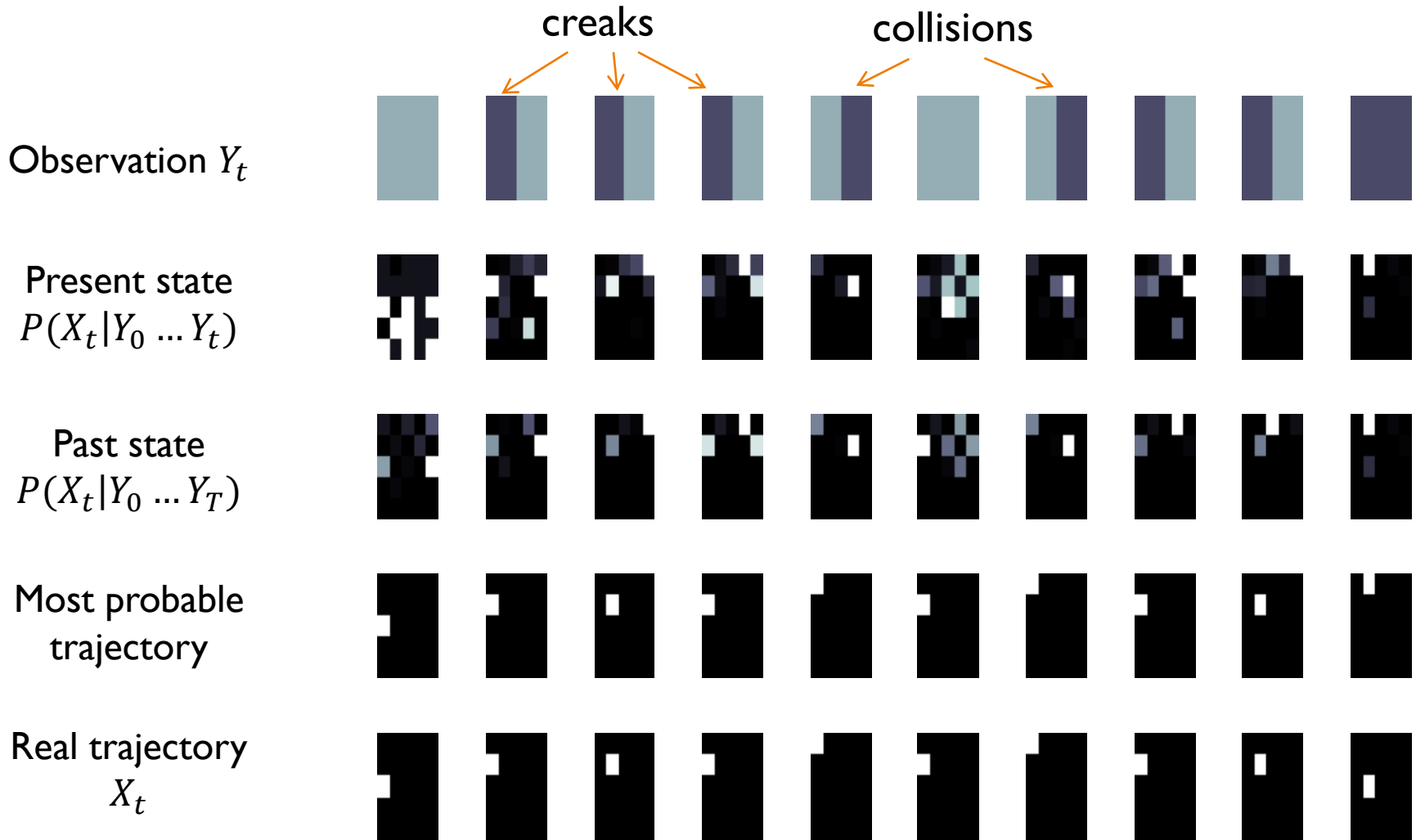


Creaking floor



Collision with an  
obstacle

# Different estimation problems with a HMM



# Partially observable process

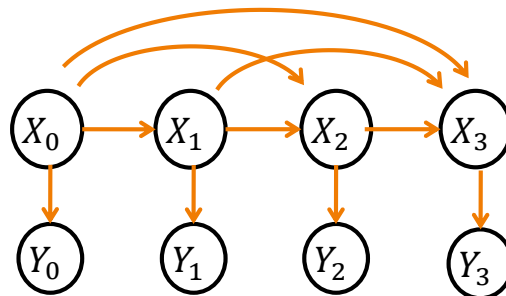
**Markov process**  $((X_0, Y_0), \dots, (X_t, Y_t), \dots)$  such that:

- State  $X_t$  is a **hidden variable**.
- State  $X_t$  is **partially observable** through **observations**  $Y_0$  à  $Y_t$
- Observation  $Y_t$  has only  $X_t$  as parent:

$$P(Y_t|X_0, \dots, X_t, \dots) = P(Y_t|X_t)$$

**Joint distribution:**

$$P(X_0, Y_0, \dots, X_t, Y_t) = P(X_0)P(Y_0|X_0) \prod_{i=1}^t P(X_i|X_0, \dots, X_{i-1})P(Y_i|X_i)$$



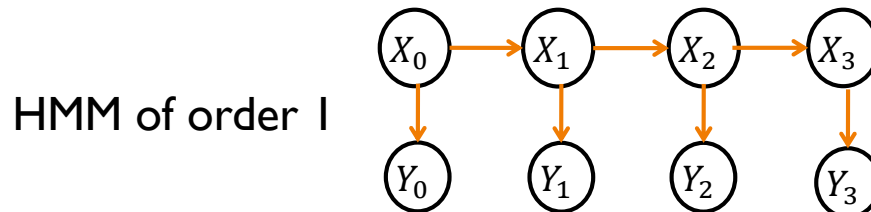
# Hidden Markov Model (HMM)

A **Hidden Markov Model** is a partially observable Markov Chain

$$P(X_0, Y_0, \dots, X_t, Y_t) = P(X_0)P(Y_0|X_0) \prod_{i=1}^t P(X_i|X_{i-1}, \dots, X_{i-k})P(Y_i|X_i)$$

A Hidden Markov Model of order  $l$  ( $k=l$ ) with discrete observation from  $l$  to  $m$  is defined by:

- A  $n \times n$  **transition matrix**:  $\mathbb{P}_t = (P(X_t = j|X_{t-1} = i))_{i,j}$
- An  $n \times m$  **emission matrix**:  $\mathbb{Q}_t = (P(Y_t = j|X_t = i))_{i,j}$





# HMM Example: tracking cachalots

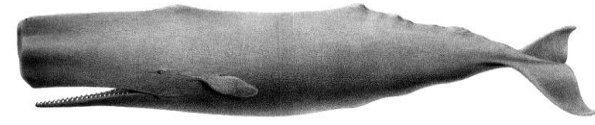


Scientifics stick a GPS device on the back of a cachalot:

- A cachalot dives in average 30 min every two hours.
- The device hibernates and wakes up few minutes every 24 hours.
- When woken up, if the device is on the sea surface, it emits its position.
- The risk for the device to become out of service is 5% per day.
- The risk for the device to come off from the cachalot is 10% per day. It then drifts on the surface.
- 50% of the messages are received by a satellite.
- 75% of the messages sent by a drifting device are received by a satellite.
- The average lifetime of the device battery is 100 days with standard deviation of 5 days.

Model the problem with a HMM as a graph and then matrices.

# Tracking cachalots: solution

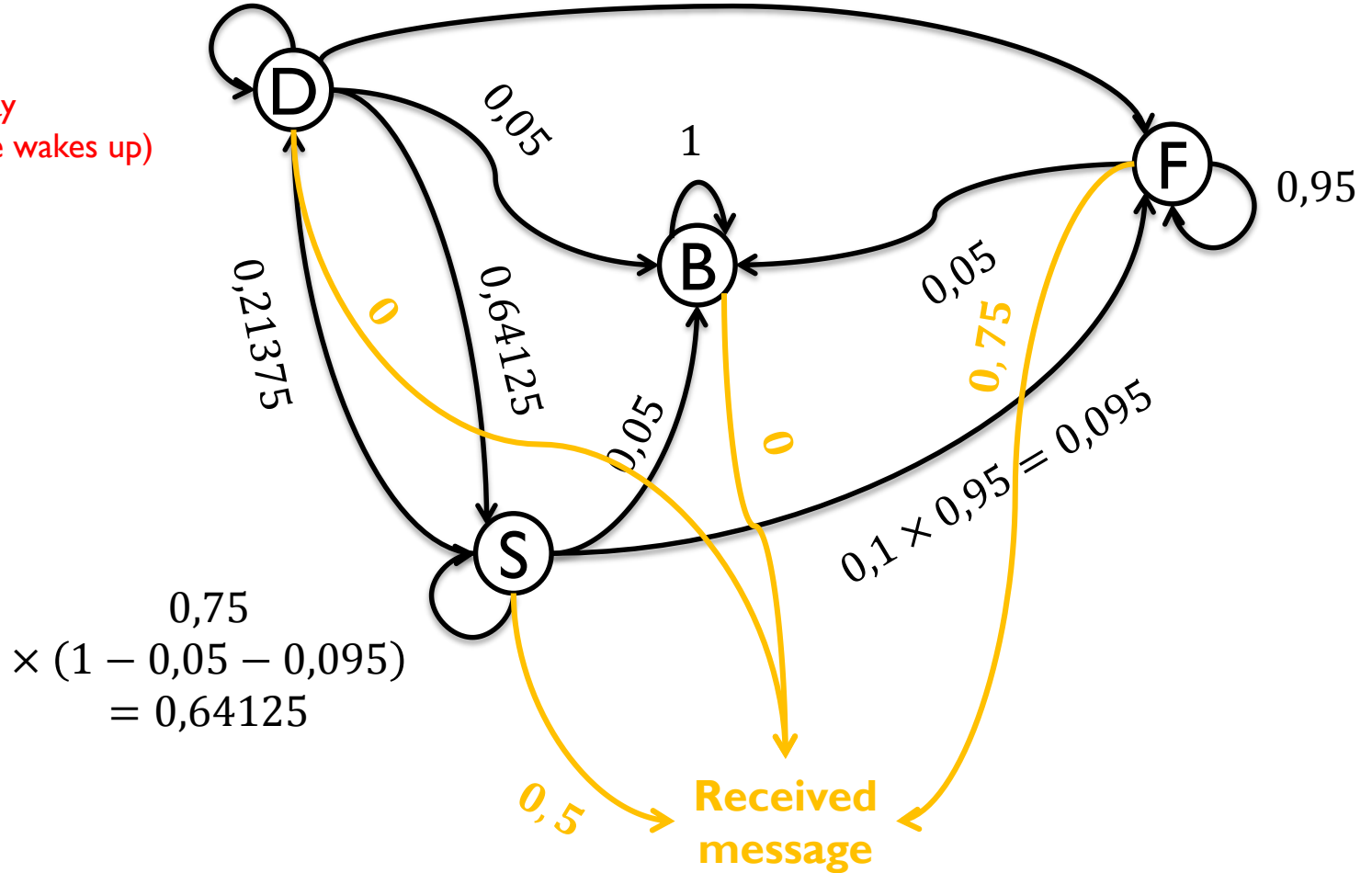


$$0,25 \times (1 - 0,05 - 0,095) = 0,21375$$

$$0,1 \times 0,95 = 0,095$$

Time step = 1 day  
(when the device wakes up)

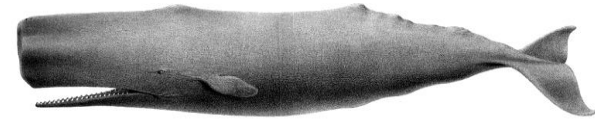
- S Surface
- D Diving
- F Floating
- B Broken



$$0,75 \times (1 - 0,05 - 0,095) = 0,64125$$



# Tracking cachalots: solution



$$\mathbb{P}_t = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 0,64125 & 0,21375 & 0,095 & 0,05 \\ 0,64125 & 0,21375 & 0,095 & 0,05 \\ 0 & 0 & 0,95 & 0,05 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$P(X_0) = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$Q_t = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} M & \neg M \\ 0,5 & 0,5 \\ 0 & 1 \\ 0,75 & 0,25 \\ 0 & 1 \end{bmatrix}$$

# Online estimation of current state (filtering)

Estimation of the present state  $P(X_t|y_0, \dots, y_t)$  from past observations :

$$P(X_t|y_0, \dots, y_t) = \frac{P(X_t, y_0, \dots, y_t)}{P(y_0, \dots, y_t)}$$
$$P(X_t|Y_0, \dots, Y_t) = \frac{\alpha_t(X_t)}{\sum_{X_t} \alpha_t(X_t)} \text{ avec } \alpha_t(X_t) = P(X_t, y_0, \dots, y_t)$$

Recursive « forward » computation of coefficients  $\alpha_t(X_t)$  :

$$\begin{aligned} \alpha_t(x) &= P(X_t = x, y_0, \dots, y_t) \\ &= \sum_{x'} P(X_t = x, X_{t-1} = x', y_0, \dots, y_t) \\ &= \sum_{x'} P(y_t|X_t = x)P(X_t = x|X_{t-1} = x')P(X_{t-1} = x', y_0, \dots, y_{t-1}) \end{aligned}$$

$$= P(y_t|X_t = x) \sum_{x'=1}^n P(X_t = x|X_{t-1} = x')\alpha_{t-1}(x')$$

$$\alpha_t(x) = (\mathbb{Q}_t)_{x,y_t} \cdot (\mathbb{P}_t^T \times \alpha_{t-1})$$

# An example of HMM : on the track of cachalots



Estimate state of device at the 4th day, after positions have been received on days 0, 2 and 3.

$$\alpha_t(x) = (\mathbb{Q}_t)_{x,y_t} \cdot (\mathbb{P}_t^T \times \alpha_{t-1})$$

$$\mathbb{P} = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 0,64125 & 0,21375 & 0,095 & 0,05 \\ 0,64125 & 0,21375 & 0,095 & 0,05 \\ 0 & 0 & 0,95 & 0,05 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$X_0 = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\mathbb{Q} = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} M & \neg M \\ 0,5 & 0,5 \\ 0 & 1 \\ 0,75 & 0,25 \\ 0 & 1 \end{bmatrix}$$



# Solution



$\otimes$  : componentwise multiplication

$$\alpha_0 = \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \otimes \begin{matrix} S \\ P \\ F \\ D \end{matrix} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0,5 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$X_0|Y_0 = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\alpha_1 = \begin{bmatrix} 0,5 \\ 1 \\ 0,25 \\ 1 \end{bmatrix} \otimes \mathbb{P}^T \times \begin{bmatrix} 0,5 \\ 0 \\ 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0,1603 \\ 0,1069 \\ 0,0119 \\ 0,0250 \end{bmatrix}$$

$$X_1|Y_{0\dots 1} = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 0,53 \\ 0,35 \\ 0,04 \\ \mathbf{0,08} \end{bmatrix}$$

$$\alpha_2 = \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \otimes \mathbb{P}^T \times \begin{bmatrix} 0,1603 \\ 0,1069 \\ 0,0119 \\ 0,0250 \end{bmatrix} = \begin{bmatrix} 0,0857 \\ 0 \\ 0,0275 \\ 0 \end{bmatrix}$$

$$X_2|Y_{0\dots 2} = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 0,76 \\ 0 \\ 0,24 \\ 0 \end{bmatrix}$$

$$\alpha_3 = \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \otimes \mathbb{P}^T \times \begin{bmatrix} 0,0857 \\ 0 \\ 0,0275 \\ 0 \end{bmatrix} = \begin{bmatrix} 0,0275 \\ 0 \\ 0,257 \\ 0 \end{bmatrix}$$

$$X_3|Y_{0\dots 3} = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 0,52 \\ 0 \\ 0,48 \\ 0 \end{bmatrix}$$

# Offline estimation of past state (smoothing)

**Estimation of  $P(X_t|y_0, \dots, y_T)$  given  $t \in \llbracket 0 \dots T \rrbracket$**

$$\begin{aligned} P(X_t = x|y_0, \dots, y_T) &\propto P(X_t = x, y_0, \dots, y_T) \\ &\propto P(Y_{t+1}, \dots, Y_T|X_t = x, y_0, \dots, y_t)P(X_t = x, y_0, \dots, y_t) \\ &\propto P(Y_{t+1}, \dots, Y_T|X_t = x) \alpha_t(x) \\ &\propto \alpha_t(x) \cdot \beta_t(x) \end{aligned}$$

$$P(X_t = x|y_0, \dots, y_T) = \frac{\alpha_t(x)\beta_t(x)}{\sum_x \alpha_t(x)\beta_t(x)} \quad \text{with} \quad \beta_t(x) = P(y_{t+1}, \dots, y_T|X_t = x)$$

# Forward/Backward algorithm

In parallel:

1. Forward recursive computation of  $\forall t, \forall x, \alpha_t(x)$

2. Backward recursive computation of  $\forall t, \forall x, \beta_t(x)$

$$\begin{aligned}\beta_t(x) &= P(y_{t+1}, \dots, y_T | X_t = x) \\ &= \sum_{x'} P(y_{t+1}, \dots, y_T, X_{t+1} = x' | X_t = x) \\ &= \sum_{x'} P(y_{t+2}, \dots, y_T | X_{t+1} = x') P(y_{t+1} | X_{t+1} = x') P(X_{t+1} = x' | X_t = x) \\ &= \sum_{x'=1}^n P(y_{t+1} | X_{t+1} = x') P(X_{t+1} = x' | X_t = x) \beta_{t+1}(x') \\ &= \boxed{\beta_t(x) = \sum_{x'=1}^n (\mathbb{Q}_t)_{x,y_t} (\mathbb{P}_{t+1})_{x,x'} \beta_{t+1}(x')} \text{ with } \forall x, \beta_T(x) = 1\end{aligned}$$

3. Compute  $P(X_t | y_0, \dots, y_T)$  from  $\alpha_t$  and  $\beta_t$



# Exemple de HMM : sur les traces des cachalots



Estimate state of device at the 4th day, after positions have been received on days 0, 2 and 3.

$$\alpha_t(x) = (\mathbb{Q}_t)_{x,y_t} \cdot (\mathbb{P}_t^T \times \alpha_{t-1}) \quad \text{and} \quad \beta_t(x) = \sum_{x'=1}^n (\mathbb{Q}_t)_{x,y_t} (\mathbb{P}_{t+1})_{x,x'} \beta_{t+1}(x')$$

$$\mathbb{P} = \begin{matrix} & S & D & F & B \\ \begin{matrix} S \\ D \\ F \\ B \end{matrix} & \begin{bmatrix} 0,64125 & 0,21375 & 0,095 & 0,05 \\ 0,64125 & 0,21375 & 0,095 & 0,05 \\ 0 & 0 & 0,95 & 0,05 \\ 0 & 0 & 0 & 1 \end{bmatrix} \end{matrix}$$

$$X_0 = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\mathbb{Q} = \begin{matrix} S \\ D \\ F \\ B \end{matrix} \begin{bmatrix} M & \neg M \\ 0,5 & 0,5 \\ 0 & 1 \\ 0,75 & 0,25 \\ 0 & 1 \end{bmatrix}$$



# Solution



$$\alpha_0 = \begin{bmatrix} 0,5 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\beta_0 = \begin{bmatrix} 0,11 \\ 0,11 \\ 0,12 \\ 0 \end{bmatrix}$$

$$\Rightarrow X_0|y_{0...3} \propto \alpha_0 \otimes \beta_0 = \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$X_0|y_0 = \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\alpha_1 = \begin{bmatrix} 0,1603 \\ 0,1069 \\ 0,0119 \\ 0,0250 \end{bmatrix}$$

$$\beta_1 = \begin{bmatrix} 0,18 \\ 0,18 \\ 0,51 \\ 0 \end{bmatrix}$$

$$\Rightarrow X_1|y_{0...3} \propto \alpha_1 \otimes \beta_1 = \begin{bmatrix} 0,53 \\ 0,35 \\ 0,12 \\ 0 \end{bmatrix}$$

$$X_1|y_{0...1} = \begin{bmatrix} 0,53 \\ 0,35 \\ 0,04 \\ 0,08 \end{bmatrix}$$

$$\alpha_2 = \begin{bmatrix} 0,0857 \\ 0 \\ 0,0275 \\ 0 \end{bmatrix}$$

$$\beta_2 = \begin{bmatrix} 0,39 \\ 0,39 \\ 0,71 \\ 0 \end{bmatrix}$$

$$\Rightarrow X_2|y_{0...3} \propto \alpha_2 \otimes \beta_2 = \begin{bmatrix} 0,63 \\ 0 \\ 0,37 \\ 0 \end{bmatrix}$$

$$X_2|y_{0...2} = \begin{bmatrix} 0,76 \\ 0 \\ 0,24 \\ 0 \end{bmatrix}$$

$$\alpha_3 = \begin{bmatrix} 0,0275 \\ 0 \\ 0,257 \\ 0 \end{bmatrix}$$

$$\beta_3 = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}$$

$$\Rightarrow X_3|y_{0...3} \propto \alpha_3 \otimes \beta_3 = \begin{bmatrix} 0,52 \\ 0 \\ 0,48 \\ 0 \end{bmatrix}$$

$$X_3|y_{0...3} = \begin{bmatrix} 0,52 \\ 0 \\ 0,48 \\ 0 \end{bmatrix}$$



# Most probable trajectory: the Viterbi algorithm

**Determine most probable sequence of state  $(x_0, \dots, x_T)$  given observations  $(y_0, \dots, y_T)$ .**

$$\operatorname{argmax}_{x_0 \dots x_T} P(x_0, \dots, x_T | y_0, \dots, y_T) \propto \operatorname{argmax}_{x_0 \dots x_T} P(x_0, \dots, x_T, y_0, \dots, y_T)$$

**Resolution by dynamic programming:**

$\mu_t(i)$ : probability of the most probable trajectory  $x_0, \dots, x_t$  such that  $x_t = i$ .

**Bellman equation:  $\mu_t(i) = \max_{x_0, \dots, x_{t-1}} P(x_0, \dots, x_{t-1}, x_t = i, y_0, \dots, y_t)$**

$t > 0$	$\rightarrow$	$\mu_t(i) = P(y_t   X_t = i) \cdot \max_j (P(X_t = i   X_{t-1} = j) \mu_{t-1}(j))$
$t = 0$	$\rightarrow$	$\mu_0(i) = P(y_0   X_0 = i) \cdot P_0(X_0 = i)$

**Final result:**

$$\max_{x_0 \dots x_T} P(x_0, \dots, x_T, y_0, \dots, y_T) = \max_i (\mu_T(i))$$



# Exemple de HMM : sur les traces des cachalots



Determine most probable trajectory at the 4th day, after positions have been received on days 0, 2 and 3.

$$\begin{aligned} t > 0 &\rightarrow \mu_t(i) = P(y_t | X_t = i) \cdot \max_j (P(X_t = i | X_{t-1} = j) \mu_{t-1}(j)) \\ t = 0 &\rightarrow \mu_0(i) = P(y_0 | X_0 = i) \cdot P_0(X_0 = i) \end{aligned}$$



Matrix reformulation where  $Diag(v)$  is diagonal matrix of coefs.  $v = [c_1, \dots, c_n]$ .

$$\begin{aligned} t > 0 &\rightarrow \mu_t = Diag(Q^{*,y_t}) \times \max_j (\mathbb{P}^T \times Diag(\mu_{t-1})) \\ t = 0 &\rightarrow \mu_0(i) = Diag(Q^{*,y_0}) \cdot \mu_0 \end{aligned}$$



# Solution



$$\mu_0 = \text{Diag} \left( \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \right) \times \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix} = \begin{matrix} S \\ P \\ F \\ D \end{matrix} \begin{bmatrix} 0,5 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\mu_1 = \text{Diag} \left( \begin{bmatrix} 0,5 \\ 1 \\ 0,25 \\ 1 \end{bmatrix} \right) \times \max_j \left( \begin{bmatrix} 0,64125 & 0,64125 & 0 & 0 \\ 0,21375 & 0,21375 & 0 & 0 \\ 0,095 & 0,095 & 0,95 & 0 \\ 0,05 & 0,05 & 0,05 & 1 \end{bmatrix} \times \text{Diag}(\mu_0) \right)$$

$$= \text{Diag} \left( \begin{bmatrix} 0,5 \\ 1 \\ 0,25 \\ 1 \end{bmatrix} \right) \times \max_j \left( \begin{bmatrix} 0,3206 & 0 & 0 & 0 \\ 0,107 & 0 & 0 & 0 \\ 0,0475 & 0 & 0 & 0 \\ 0,025 & 0 & 0 & 0 \end{bmatrix} \right) = \begin{bmatrix} 0,160 \\ 0,107 \\ 0,012 \\ 0,025 \end{bmatrix}$$



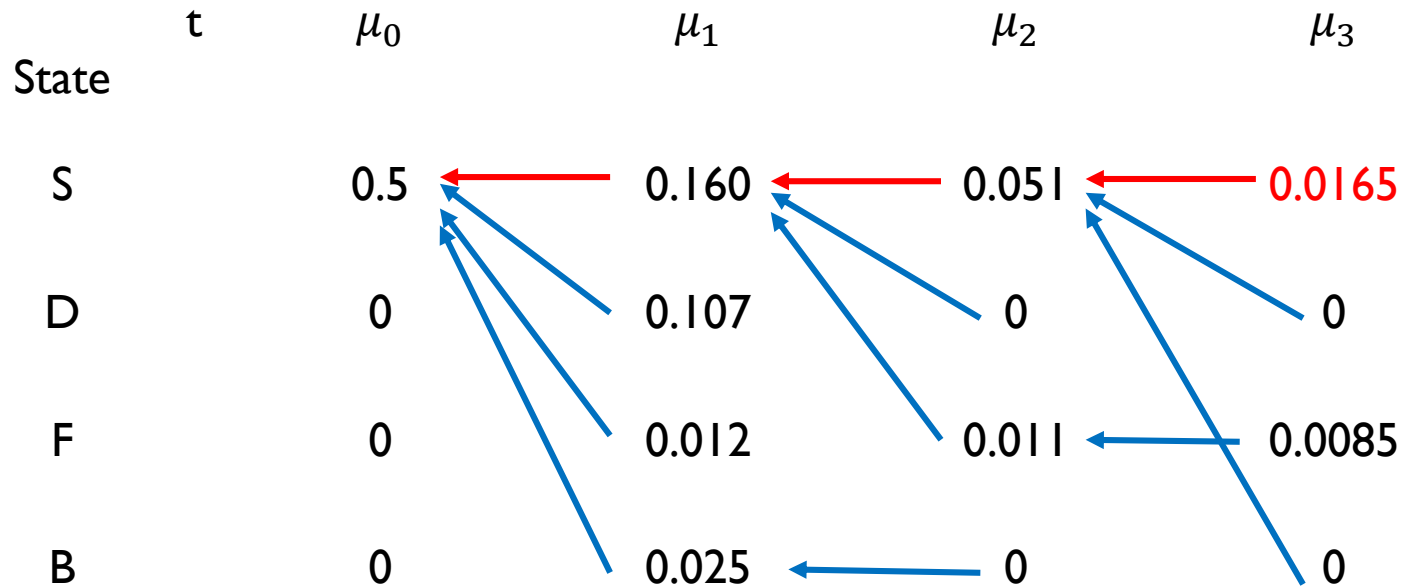
# Solution



$$\begin{aligned}
\mu_2 &= \text{Diag} \left( \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \right) \times \max_j \left( \begin{bmatrix} 0,64125 & 0,64125 & 0 & 0 \\ 0,21375 & 0,21375 & 0 & 0 \\ 0,095 & 0,095 & 0,95 & 0 \\ 0,05 & 0,05 & 0,05 & 1 \end{bmatrix} \times \text{Diag}(\mu_1) \right) \\
&= \text{Diag} \left( \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \right) \times \max_j \left( \begin{bmatrix} \mathbf{0,1028} & 0,068 & 0 & 0 \\ \mathbf{0,034} & 0,023 & 0 & 0 \\ \mathbf{0,152} & 0,010 & 0,0113 & 0 \\ 0,008 & 0,005 & 0,0006 & \mathbf{0,025} \end{bmatrix} \right) = \begin{bmatrix} 0,0514 \\ 0 \\ 0,0114 \\ 0 \end{bmatrix} \\
\mu_3 &= \text{Diag} \left( \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \right) \times \max_j \left( \begin{bmatrix} 0,64125 & 0,64125 & 0 & 0 \\ 0,21375 & 0,21375 & 0 & 0 \\ 0,095 & 0,095 & 0,95 & 0 \\ 0,05 & 0,05 & 0,05 & 1 \end{bmatrix} \times \text{Diag}(\mu_2) \right) \\
&= \text{Diag} \left( \begin{bmatrix} 0,5 \\ 0 \\ 0,75 \\ 0 \end{bmatrix} \right) \times \max_j \left( \begin{bmatrix} \mathbf{0,033} & 0 & 0 & 0 \\ \mathbf{0,011} & 0 & 0 & 0 \\ 0,005 & 0 & \mathbf{0,011} & 0 \\ \mathbf{0,003} & 0 & 0,0006 & 0 \end{bmatrix} \right) = \begin{bmatrix} 0,0165 \\ 0 \\ 0,0081 \\ 0 \end{bmatrix}
\end{aligned}$$



# Solution



→ Most probable trajectory :  $(x_0, x_1, x_2, x_3) = (S, S, S, S)$

# Learning a HMM:

**Problem:** learn  $\theta = (P_0, (\mathbb{P}_t, \mathbb{Q}_t)_{t \geq 0})$  from  $N$  i.i.d sequences  
 $(y_0^1 \dots y_{T_1}^1) \dots (y_0^N \dots y_{T_N}^N)$

## Two cases:

- States  $(x_0^1 \dots x_{T_1}^1) \dots (x_0^N \dots x_{T_N}^N)$  are known (expert annotation)
  - Split problem in two parts:
    1. Learn Markov chain  $(P_0, (\mathbb{P}_t)_{t \geq 0})$
    2. Learn emission matrices  $(\mathbb{Q}_t)_{t \geq 0}$  using MLE
- States  $(x_0^1 \dots x_{T_1}^1) \dots (x_0^N \dots x_{T_N}^N)$  are unknown
  - EM must be used to learn distribution of hidden states
  - Baum Welch algorithm

# The Baum-Welch algorithm

1. Initialize randomly  $\theta = (P_0, (\mathbb{P}_t, \mathbb{Q}_t)_{t \geq 0})$  for  $n$  fixed
2. **E-step:** estimate  $(a_t^i)_{i,t}$  and  $(B_t^i)_{i,t}$  from  $\theta$  using backward-forward algorithm

- $\forall i, \forall t \in \{0, \dots, T_i\}, \forall s$ , distribution  $a_t^i$  of  $X_t^i$

$$a_t^{i,s} \leftarrow P(X_t^i = s \mid y_0^i \dots y_{T_i}^i, \theta) \propto \alpha_t(s) \beta_t(s)$$

- $\forall i, \forall t \in \{0, \dots, T_i\}, \forall s, \forall s'$  distribution  $B_t^i$  of transition  $X_t^i \rightarrow X_{t+1}^i$ :

$$B_t^{i,s,s'} \leftarrow P(X_t^i = s, X_{t+1}^i = s' \mid y_0^i \dots y_{T_i}^i, \theta) \propto \alpha_t(s) \mathbb{P}_t^{s,s'} \mathbb{Q}_t^{s',y_{t+1}^i} \beta_t(s')$$

3. **M-step:** learn  $(P_0, (\mathbb{P}_t, \mathbb{Q}_t)_{t \geq 0})$  from  $(a_t^i)_{i,t}$  and  $(B_t^i)_{i,t}$

- $\forall t, \forall s, \forall y$

$$P_0^s \propto \sum_i a_0^{i,s} \quad \text{and} \quad \mathbb{Q}_t^{s,y} \propto \sum_i a_t^{i,s} \mathbb{1}(y_t^i = y)$$

- $\forall t, \forall s, \forall s'$ ,

$$\mathbb{P}_t^{s,s'} \propto \sum_i B_t^{i,s,s'}$$

4. Go to step 2 until convergence of  $\theta$